## Appendix: Derivation of the Distribution Truncation in EKF manner

The problem is to calculate the new mean  $oldsymbol{q}_k$  and covariance  $oldsymbol{Q}_k$  of the region which satisfies an inequality constraint  $\varphi_k^T x \leq b_k$  under the prior normal distribution with the mean  $q_{k-1}$  and covariance  $Q_{k-1}$ . It is reduced to a simpler case that the prior mean is 0, the prior covariance is I and the constraint is  $x'_1 \leq c_k$ , by applying the following linear transform:

$$x' = RW^{-\frac{1}{2}}T^{T}(x - q_{k-1})$$

$$\tag{1}$$

where R and T are othogonal matrices and W is diagonal matrix respectively, and they satisfy the following equations:

$$TWT^T = Q_{k-1} \tag{2}$$

$$\mathbf{R}\mathbf{W}^{\frac{1}{2}}\mathbf{T}^{T}\varphi_{k} = ((\varphi_{k}^{T}\mathbf{Q}_{k-1}\varphi_{k})^{\frac{1}{2}}, 0, \cdots, 0)^{T}$$

$$(3)$$

$$c_k = \frac{b_k - \varphi_k^T \mathbf{q}_{k-1}}{\left(\varphi_k^T \mathbf{Q}_{k-1} \varphi_k\right)^{\frac{1}{2}}}.$$
(4)

Since the transformed distribution is a isotropic (note that its covariance is an identity matrix), the components are independent to each other and the distribution can be represented as the product of the marginal distributions.

$$p(x') = \prod_{i} p_i(x'_i) = \prod_{i} \frac{1}{\sqrt{2\pi}} \exp(-\frac{{x'_i}^2}{2})$$
 (5)

Here, the inequality constraint is  $x'_1 \leq c_k$ , namely the probability density in the region  $x'_1 > c_k$ should be zero by the distribution truncation. Therefore, the original distribution should be normalized so that the definite integral of the normalized distribution over  $x_1 \leq c_k$  is 1.

$$\int_{-\infty}^{c_k} \cdots \int_{-\infty}^{\infty} p(\mathbf{x}') dx_1' \cdots dx_n' = \int_{-\infty}^{c_k} \frac{1}{\sqrt{2\pi}} \exp(-x_1'^2) dx_1'$$
$$= \frac{1}{2} \left\{ 1 + \operatorname{erf}\left(\frac{c_k}{\sqrt{2}}\right) \right\}$$
(6)

where  $erf(\cdot)$  denotes the error function:

$$\operatorname{erf}(x) = \int_{-\infty}^{x} \frac{1}{\sqrt{2\pi}} \exp(-\frac{t^2}{2}) dt. \tag{7}$$

This function is not an analytical function but it can be quite rapidly calculated using a common numerical calculation package. The normalized distribution is as follows:

$$p'(\mathbf{x'}) = \alpha \prod_{i} p_i(x_i') \tag{8}$$

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$$\alpha \equiv \frac{\sqrt{2}}{\sqrt{\pi}(1 + \operatorname{erf}(c_k/\sqrt{2})}.$$
(8)

The mean and covariance of the above distribution are obtained as follows, using the fact that each  $x_i'$  is independent:

$$\mu_{1} = E[x'_{1}] = \alpha \int_{-\infty}^{c_{k}} x'_{1} \exp(-x'_{1}^{2}/2) dx'_{1}$$
$$= -\alpha \exp(-\frac{c_{k}^{2}}{2})$$
(10)

$$\mu_i = E[x_i'] = 0 \ (i \neq 1)$$
 (11)

$$\sigma_{11}^2 = E[x_1' - \mu_1][x_1' - \mu_1] = \alpha \int_{-\infty}^{c_k} (x_1' - \mu_1)^2 \exp(-\frac{{x_1'}^2}{2}) dx_1'$$

$$= 1 - \alpha c_k \exp(-\frac{c_k^2}{2}) - \mu_1^2 \tag{12}$$

$$\sigma_{ii}^2 = E[x_i' - \mu_i][x_i' - \mu_i] = 1 \ (i \neq 1)$$
(13)

$$\sigma_{ij}^2 = E[x_i' - \mu_i][x_j' - \mu_j] = 0 \ (i \neq j). \tag{14}$$

By rewritten the above equations introducing a variable

$$\nu_k \equiv \mu_1 \tag{15}$$

$$= -\sqrt{\frac{2}{\pi}} \exp(-\frac{c_k^2}{2})/(1 + \operatorname{erf}(\frac{c_k}{\sqrt{2}})), \tag{16}$$

the mean and covariance are

$$\boldsymbol{\mu}_k = (\nu_k, 0, \cdots, 0) \tag{17}$$

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$$S_k = \text{diag}\{1 + c_k \nu_k - \nu_k^2, 1, \dots, 1\}$$
(17)

where  $\operatorname{diag}\{a,b,\cdots\}$  represents a diagonal matrix whose diagonal elements are  $a,b,\cdots$ . Then the truncated mean and variance are expressed as

$$\boldsymbol{q}_k = \boldsymbol{T} \boldsymbol{W}^{\frac{1}{2}} \boldsymbol{R}^T \boldsymbol{\mu}_k + \boldsymbol{q}_{k-1} \tag{19}$$

$$Q_k = TW^{\frac{1}{2}}R^TS_kRW^{\frac{1}{2}}T^T$$
 (20)

by applying inversely the transform Eq.1.